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Labour market, health and well-being: economic analysis using panel data

Plan for today

- Other types of fixed effects
- Random effects model
- Test and diagnostics: RE or FE?
- Correlated Random Effects
- Limitations of panel data estimation
- How to interpret your results and present possible limitations

Other types of fixed effects

- Time FE (use time dummies to understand if there is a specific effect of a certain point in time)
- Country/Region FE (use country or region dummies to control for the time-invariant characteristics of a certain geographical area)

Other types of fixed effects

- School/Class FE (to capture the common factors in a school/class, eg teachers' attitudes, environment, prejudices etc)
- Sibling FE (to capture common characteristics of kids raised in the same family)

Time FE

- To see if time FE (eg year dummies) are needed we may use testparm after running the command in Stata
- This is a joint test to see if the dummies for all years are equal to 0
- `xtreg sclfsato age_dv i.wave i.sex_dv married separated widow i.hiqal_dv self_employed unemployed outlforce, fe`

```
. testparm i.wave
```

```
( 1)  2.wave = 0  
( 2)  3.wave = 0  
( 3)  4.wave = 0  
( 4)  5.wave = 0  
( 5)  6.wave = 0  
( 6)  7.wave = 0  
( 7)  8.wave = 0  
( 8)  9.wave = 0
```

```
F( 8,140146) = 95.68  
Prob > F = 0.0000
```



We reject the null hypothesis that the coefficients for all years are jointly equal to 0 and so we need time FE

Random effects estimator

- **FE estimator wipes out the between variation**
(variation across cross sections)
- If we believe that the a_i 's are not correlated with the regressors, we should resort to more efficient estimators
- RE estimator could be a feasible estimator

Random effects estimator

$$Y_{it} = \alpha + x_{it}\beta + a_i + \varepsilon_{it}$$

- Suppose a_i is purely random with mean 0 and variance σ_a^2
- In other words, we are assuming a_i is **uncorrelated** with the regressors
- This is a **very strong assumption** (even stronger than what needed for FE)

Random effects estimator

- The RE estimator is a **weighted average of a within and between model**

$$(y_{it} - \bar{y}_i) + (1 - \hat{\theta}_i)\bar{y}_i = (x_{it} - \bar{x}_i)\beta + (\epsilon_{it} - \bar{\epsilon}_i) + (1 - \hat{\theta}_i)(\alpha + \bar{x}_i\beta + a_i + \bar{\epsilon}_i)$$

- **In Stata, we use xtreg, re**

Random effects estimator

$$Y_{it} = \alpha + x_{it}\beta + a_i + \varepsilon_{it}$$

- The logic of RE is that, unlike FE, the variation across individuals is assumed to be random and uncorrelated with the independent variables in the model
- “...the crucial distinction between fixed and random effects is **whether the unobserved individual effect embodies elements that are correlated with the regressors** in the model, not whether these effects are stochastic or not” [Green, 2008, p.183]

Random effects estimator

$$Y_{it} = \alpha + x_{it}\beta + a_i + \varepsilon_{it}$$

- An advantage of RE is that **we can include time-constant independent variables** (eg gender and in many cases education)
- In the RE model, you need to include individual characteristics that may or may not influence the predictor variables.
- However, some variables may not be available
- Omitted variable bias in the model?

Comparing RE and FE

- Let's use the last example we made and compare RE and FE estimates
- Use longitudinal_td.dta
- `xtreg sclfsato age_dv i.sex_dv married separated widow i.hiqal_dv self_employed unemployed outforce, fe`
- `estimates store FE`
- `xtreg sclfsato age_dv i.sex_dv married separated widow i.hiqal_dv self_employed unemployed outforce, re`
- `estimates store RE`
- `estimates table FE RE, b se`

Comparing RE and FE

. estimates table FE RE, b se

Variable	FE	RE
age_dv	-.00642314 .00122866	.00616264 .00041506
sex_dv		
Female	(omitted)	.0415845 .01256114
married	.16085272 .02412544	.22832874 .01476111
separated	-.02995826 .03080957	-.20300746 .02036037
widow	-.04190701 .04028489	.07657132 .02542923
hiqual_dv		
Other hig..	.1084236 .05597975	-.06140337 .02010287
A level etc	.17343366 .04605859	-.07522068 .01781304
GCSE etc	.18094291 .06000858	-.1816244 .01796659
Other qual	.08059962 .08320697	-.22135884 .02208995
No qual	.08769546 .088886	-.25807731 .0206778
self_emplo~d	.02593077 .01978107	.00009619 .01615603
unemployed	-.21119275 .02047336	-.34781119 .01791513
outlforce	.02052477 .013382	.00540641 .01055973
_cons	5.3308938 .08000153	4.838006 .02384111

legend: b/se

RE or FE

- If the individual FE is correlated with the regressors, we must use FE (the only consistent estimator)
- If individual FE is uncorrelated with the regressors, both RE and FE are consistent but RE is more efficient
- This is because RE uses between and within variation

RE or FE

- **The Hausman test can be used to choose RE or FE**
- It starts with the idea that both estimators are consistent under the null hypothesis of no correlation
- Under the null hypothesis the estimates of β should not differ systematically
- Use **hausman** in Stata to run the test

RE or FE

- The intuition is: use RE only if Hausman test says it's ok!
- FE is consistent. If RE does not differ too much then we can use RE
- If you **reject H_0** , use FE
- If you are **not** able to reject H_0 , use RE

RE or FE

hausman FE RE, sigmamore

	Coefficients			
	(b) FE	(B) RE	(b-B) Difference	sqrt(diag(V_b-V_B)) S.E.
age_dv	-.0064231	.0061626	-.0125858	.0011554
married	.1608527	.2283287	-.067476	.0190593
separated	-.0299583	-.2030075	.1730492	.0230919
widow	-.041907	.0765713	-.1184783	.0312049
hiqual_dv				
2	.1084236	-.0614034	.169827	.0521998
3	.1734337	-.0752207	.2486543	.0424364
4	.1809429	-.1816244	.3625673	.0572078
5	.0805996	-.2213588	.3019585	.0801552
9	.0876955	-.2580773	.3457728	.0863776
elf_emplo~d	.0259308	.0000962	.0258346	.0113875
unemployed	-.2111927	-.3478112	.1366184	.0098776
outlforce	.0205248	.0054064	.0151184	.0082034

b = consistent under Ho and Ha; obtained from xtreg

B = inconsistent under Ha, efficient under Ho; obtained from xtreg

Test: Ho: difference in coefficients not systematic

$$\begin{aligned}\text{chi2(12)} &= (b-B)' [(V_b-V_B)^{-1}] (b-B) \\ &= 606.75 \\ \text{Prob>chi2} &= 0.0000\end{aligned}$$



If this is <0.05 use FE

RE or FE

- Given the test, we reject the null hypothesis of no correlation
- **We conclude that the estimators are different and we must use FE**

Limitations of panel data analysis

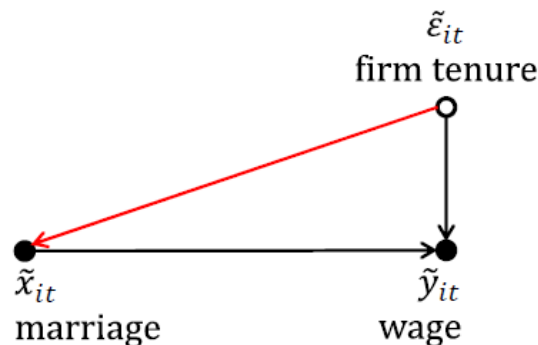
- **FE estimates are biased under endogeneity**

$$E(x_{is} \varepsilon_{it}) \neq 0 \text{ for } s, t=1, \dots, T$$

- Endogeneity can have several sources
 - Unobserved time-varying confounders
 - Measurement error (errors in reporting x)
 - Y affects X (reverse causality)
 - Endogenous selection bias
- When using FE, we need to argue that none of these sources of endogeneity is problematic

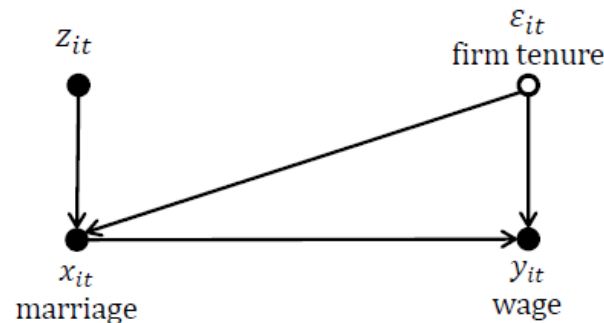
Sources of endogeneity

- **Unobserved time-varying variables**
 - eg: is there a marriage premium (does marriage increase wages?)
 - Strict exogeneity assumption: firm tenure does not affect chances of marriage
 - However: longer firm tenure may indicate a more secure job and this in turn may affect chances of marriage



Sources of endogeneity

- How to solve this problem?
 - Instrumental Variables can be used
 - Stata routine is `xtivreg`
 - In this example, we need to find a variable (instrument) that affects marriage but does not directly affect wage



Sources of endogeneity

- **Measurement error**

- If the model includes only one variable, measurement error generally produce “**attenuation bias**”
- With more variables, the direction of the bias is unknown
- When using FD or FE estimator, the measurement error problem may be amplified
- Taking the difference of two unreliable measures can produce an even more unreliable measure
- However, FD or FE reduces the unobserved heterogeneity problem and simulation studies show this bias dominates

Sources of endogeneity

- Measurement error
 - People respond to the same question in different ways
 - Example: life satisfaction. Different people mean something different when they say they are completely satisfied with their life
 - The problem is amplified with ranking scales (eg rank your life satisfaction from 1 to 7)
 - If this type of error is related to time-invariant characteristics (eg personality traits) FE estimates will not be biased

Sources of endogeneity

- Reverse causality
 - Reverse causality (Y affects X)
 - Reciprocal causality (Y affects X and X affects Y)
- **Any econometric model is biased under reciprocal causality!**
- We can mitigate the risk of reverse causality by using panel data and looking at the timing of the events

Sources of endogeneity

- Selection into treatment
 - Treatment X affects outcome Y
 - People with certain levels of Y are more likely to get X
 - This means some people may select into treatment
 - FE can help to deal with selection
 - Example: people living with a partner from a different ethnic group or religion are less likely to have stereotypes against immigrants
 - BUT
 - People with less stereotypes are more likely to marry someone with a different background
 - FE can help in this context (even if it may not be perfect)

Panel attrition

- Panel attrition can be a threat to panel analysis
- However, the main problem comes from attrition deriving from unobservable variables
- Attrition deriving from time-invariant variables is not a problem for FE estimation
- Example: people with certain personality traits are more likely to leave the survey

Correlation over time

- Panel variables can be correlated over time
 - Outcomes depend on characteristics that are fixed over time
 - FE estimates are unbiased
 - Past outcomes affect future outcomes
 - We can use models with lagged dependent variable as a regressor (dynamic panel models)
 - However, FE is not feasible here (strict exogeneity assumption is violated)
 - We can use IV estimation

How to interpret your results – Caution!

- No estimation method will ever be perfect!
- If you have doubts about the strict exogeneity assumption, **explain the *caveats* carefully after your present the results**
- Use caution in interpreting the results as causal if you have doubts about the assumption
- It is better to show caution than to look overconfident!

How to interpret your results – Caution!

- **It is very hard to rule out reverse causality by simply using panel data**
- The timing of events is not necessarily helpful, because there may be anticipation effects and measurement error
- Sometimes it is hard to know whether X or Y changed first
- **Be upfront and explain the potential limitations of your work!**